# Estimation of Galaxy Luminosity Distributions from Incomplete X-ray and Optical Survey Data

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# Motivation: log(N) - log(S)



### Luminosity function

Goal: Estimate X-ray source intensities and obtain luminosity function.

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### A New Strategy for a Key Challenge



#### Sketch: Source population (detection cases)

- Challenge: Incompleteness in X-ray data
- Low-intensity sources most prevalent, but least likely detected in X-ray
- Strategy: Leverage optical data, combining X-ray and optical surveys
- Account for missing sources via incompleteness functions

## X-ray and Optical Incompleteness

#### **Sketch: Incompleteness functions**



### Additional Complications:



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- Background contamination, exposure and off-axis corrections
- Overlapping optical sources (blue)
- X-ray source detection (yellow), and optical matching

### Source Model

 In source region *i*, observed photon counts Y<sub>i</sub> are sum of latent background B<sub>i</sub> and source S<sub>i</sub>

$$Y_j = \mathcal{S}_j + \mathcal{B}_j. \tag{1}$$

Arrival of photons at detector as Poisson process, with source model

$$S_i | \lambda_i \overset{\text{indep}}{\sim} \text{Poisson}(e_i \lambda_i \mathcal{T})$$
 (2)

• Background count  $\mathcal{B}_i$  with rate  $\xi$  in source region *i* modeled via

$$\mathcal{B}_i | \xi \overset{\text{indep}}{\sim} \mathsf{Poisson}(a_i \xi \mathcal{T}).$$
 (3)

Data	Description
ai	area of the source region
Yi	counts collected in source region (of area $a_i$ )
ei	telescope effective area [cm <sup>2</sup> ] at source location
Α	area of the background region
X	collected background counts

### **Background Model**

- Background rate  $\xi$  (count/s/pixel) **uniform** across source regions.
- Observed photon count in background modeled via

$$X|\xi \sim \text{Poisson}(A\xi\mathcal{T}),$$
 (4)



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# Estimation of the Luminosity Function



- Piece-wise linear
   log(N) log(S) relation
   assumed.
- Breakpoints τ<sub>1</sub>, τ<sub>2</sub>, slopes θ<sub>1</sub>, θ<sub>2</sub>

•  $(\lambda_1, \ldots, \lambda_n)$  independent with double-Pareto population

$$f(\lambda_i|\theta_1,\theta_2,\tau_1,\tau_2) = \left(\frac{\theta_1}{\tau_1}\right) \left(\frac{\lambda_i}{\tau_1}\right)^{-(\theta_1+1)} \mathbf{1}_{\{\tau_1 \le \lambda_i \le \tau_2\}} + \left(\frac{\tau_2}{\tau_1}\right)^{-\theta_1} \left(\frac{\theta_2}{\tau_2}\right) \left(\frac{\lambda_i}{\tau_2}\right)^{-(\theta_2+1)} \mathbf{1}_{\{\tau_2 \le \lambda_i \le \infty\}}$$

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### **Incompleteness Correction**

• For each source in population, introduce indicator Z<sub>i</sub>, with

$$Z_{i} = \begin{cases} 0 & \text{if source } i \text{ not observed} \\ 1 & i \text{ observed in optical only} \\ 2 & i \text{ observed in both} \end{cases} \begin{pmatrix} \text{w/p } 1 - g(\lambda_{i}) - h(\lambda_{i}|\phi) \\ (\text{w/p } h(\lambda_{i}|\phi)) \\ (\text{w/p } g(\lambda_{i}), \text{ with } g \text{ known}) \end{cases}$$

• Joint distributions of Z and  $\lambda$  via

$$p(\lambda, Z = 2|\boldsymbol{\theta}) = f(\lambda; \boldsymbol{\theta})g(\lambda)$$
(6)

$$p(\lambda, Z = 1 | \boldsymbol{\theta}, \phi) = f(\lambda; \boldsymbol{\theta}) h(\lambda | \phi)$$
(7)

• Binning based on  $\lambda$ . Let k = 1, ..., K be bins on  $\lambda$ , with bin size  $\Delta_{\lambda_k}$ 

 $X_1^{(k)} = #$  of sources in bin k with Z = 1, with  $X_1^{(k)} \sim Pois(f(\lambda_k; \theta) \Delta_{\lambda_k} h(\lambda_k | \phi) \psi)$ 

 $X_2^{(k)} = #$  of sources in bin k with Z = 2, with  $X_2^{(k)} \sim Pois(f(\lambda_k; \theta) \Delta_{\lambda_k} g(\lambda_k) \psi)$ 

(5)

### **Computational Details:**

### Incompleteness function:

Incompleteness function dependend on nuisance parameters:
 L (off-axis angle), e (effective area), ξ (background rate)

$$p(\lambda, L, \xi, e, Z = 2|\theta) = p(\lambda, L, \xi, e|\theta)P(Z = 2|\lambda, L, \xi, e)$$
(8)

• Assuming independence of  $\lambda$  and  $L, \xi, e$ , this yields for  $\lambda$ 

$$p(\lambda, Z = 2|\boldsymbol{\theta}) = f(\lambda|\boldsymbol{\theta}) \int p(L)p(\xi)p(\boldsymbol{\theta})P(Z = 2|\lambda, L, \xi)dLd\xi d\boldsymbol{\theta}$$

### Metropolis-within-Gibbs sampler:

Updating source/background model and incompleteness via

1.) 
$$(\xi, \boldsymbol{\lambda})^{(t)} \sim p(\xi, \boldsymbol{\lambda} | \boldsymbol{\theta}^{(t-1)}, \boldsymbol{\psi}^{(t-1)}, \boldsymbol{\phi}^{(t-1)}, \boldsymbol{X}, \boldsymbol{Y}, N_{\text{obs}})$$
 (9)

2.) 
$$(\boldsymbol{\theta}, \boldsymbol{\psi}, \boldsymbol{\phi})^{(t)} \sim \boldsymbol{p}(\boldsymbol{\theta}, \boldsymbol{\psi}, \boldsymbol{\phi} | \boldsymbol{\lambda}^{(t)}, \boldsymbol{Z}_{\text{obs}}),$$
 (10)

- Simulate realistic X-ray sources, mimicking data from the Chandra Deep Field South (CDFS)
- Assuming optical and X-ray survey incompleteness (resembling previous studies (Stampoulis 2018))
- Expected population size  $\psi = 1000$ , with only  $N_{obs} = 124$  observed (in either both, or only optical survey)

### Simulation Study – Results:

### Marginal posterior distributions (parameters of interest):



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### Simulation Study – Results:



- Correlation between τ<sub>1</sub> (population minimum) and θ<sub>1</sub> (first power-law slope)
- Higher τ<sub>1</sub> leads to steeper log(N) log(S)
- Lower τ<sub>1</sub> leads to strong overestimation of population size, and flat log(N) – log(S)

### Data Analysis: Chandra Deep Field South (CDFS)

- 1524 galaxies observed (without overlap) in subregion of CDFS
- 70 galaxies observed in X-ray and optical, 1454 in optical only
- Source counts extracted at optical locations



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### Data Analysis: Chandra Deep Field South



• Expected population size  $\psi = 110,704$  (mean posterior).

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# **Ongoing Work: Overlapping Sources**



• Photon counts  $Y_{I(s)}$  modelled in each of 14 segments *s* of overlap.

- For highlighted segment:  $I(s) = \{1, 2, 4\}$  with counts  $Y_{I(s)}$ .
- Y<sub>I(s)</sub> consists of mixture of photons from sources in s and background

$$Y_{\mathcal{I}(s)} = \sum_{i \in \mathcal{I}(s)} S_{s,i} + \mathcal{B}_{\mathcal{I}(s)}.$$
 (11)

# **Concluding Remarks**

- Estimation of X-ray luminosity distributions is challenging!
  - Strong incompleteness in the high density region of the population
- Optical surveys can be leveraged, complementing sources in the incomplete, low-intensity end
  - Principled incompleteness corrections via hierarchical Bayes model

- Further challenges/extensions:
  - Incorporating overlapping sources
  - Improving survey matching
  - Uncertainty on source locations
  - Zero-inflated population model

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- Udaltsova, I. S. (2014). The Universe at Your Fingertips: Bayesian Modeling and Computation in Problems of Observational Cosmology. University of California, Davis.
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### Thank you very much for your time!

### Chandra Deep Field Survey South – Merged Event File



### Additional Data available:

Data	Description
ai	area of the source region
Y <sub>i</sub>	counts collected in source region (of area $a_i$ )
di	area of the background region (around source <i>i</i> )
Xi	background counts collected in source region (of area $d_i$ )
ei	telescope effective area [cm <sup>2</sup> ]
ri	proportion of photons expected to fall in source region $\equiv 1$
bg-sur-bri	background counts /pixel (for incompleteness correction)
off-axis	(off-axis angle - needed for the incompleteness correction)
sign	(source S/N ratio)

- Data available from the Chandra Deep Field Catalogue.
- *n* = 358 X-ray sources detected in data set.
- Observation time  $\mathcal{T}$  in seconds ( $\mathcal{T} = 1960631$ ).
- The count-rate to flux conversion for the reference point is 1.06E-11 erg/s/cm<sup>2</sup>/cnt/s

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### Likelihood Functions of Source Model

### Background case 1:

• Likelihood function for  $(\xi, \lambda)$ , with  $\lambda = (\lambda_1, \dots, \lambda_n)$ ,

$$L(\xi, \boldsymbol{\lambda} | \mathbf{D}) = \exp(-A\mathcal{T}\xi) \frac{(A\mathcal{T}\xi)^{X}}{X!} \prod_{i=1}^{n} \exp[-(a_{i}\xi + r_{i}e_{i}\lambda_{i})\mathcal{T}] \frac{[(a_{i}\xi + r_{i}e_{i}\lambda_{i})\mathcal{T}]^{Y_{i}}}{Y_{i}!}$$

### Background case 2:

• Likelihood function for  $(\boldsymbol{\xi}, \boldsymbol{\lambda})$ , with  $\boldsymbol{\lambda} = (\lambda_1, \dots, \lambda_n)$  and  $\boldsymbol{\xi} = (\xi_1, \dots, \xi_n)$  $L(\boldsymbol{\xi}, \boldsymbol{\lambda} | \mathbf{D}) = \prod_{i=1}^n \exp(-a_i \mathcal{T} \xi_i) \frac{(a_i \mathcal{T} \xi_i)^{X_i}}{X_i!} \prod_{i=1}^n \exp[-(a_i \xi + r_i e_i \lambda_i) \mathcal{T}] \frac{[(a_i \xi + r_i e_i \lambda_i) \mathcal{T}]^{Y_i}}{Y_i!}$ 

### Likelihood and Posterior

- Assume that  $X_1^{(k)}, X_2^{(k)} \perp X_1^{(l)}, X_2^{(l)} | \theta, \phi, \psi$ , for all  $k \neq l$ , with k, l = 1, ..., K.
- Log-likelihood (Assuming *λ* (observed sources) known):

$$\begin{split} \ell(\boldsymbol{\theta}, \phi, \psi | \{X_1^{(k)}\}_{k=1}^K, \{X_2^{(k)}\}_{k=1}^K) &= -\psi \sum_k f(\lambda^{(k)}; \boldsymbol{\theta}) \Delta_{\lambda_k} [h(\lambda^{(k)} | \phi) + g(\lambda^{(k)})] \\ &+ \sum_{\{k: X_1^{(k)} > 0\}} X_1^{(k)} \log(f(\lambda^{(k)}; \boldsymbol{\theta})) + \sum_{\{k: X_2^{(k)} > 0\}} X_2^{(k)} \log(f(\lambda^{(k)}; \boldsymbol{\theta})) \\ &+ \sum_{\{k: X_1^{(k)} > 0\}} X_1^{(k)} \log(h(\lambda^{(k)} | \phi)) + \sum_{\{k: X_2^{(k)} > 0\}} X_2^{(k)} \log(g(\lambda^{(k)})) \\ &+ N_{obs} \log(\psi) \end{split}$$

Full posterior distribution:

 $p(\xi, \boldsymbol{\lambda}, \boldsymbol{\theta}, \boldsymbol{\phi}, \boldsymbol{\psi} | \boldsymbol{Y}, \boldsymbol{X}, Z_{\text{obs}}) \approx p(\boldsymbol{Y}, \boldsymbol{X} | \boldsymbol{\xi}, \boldsymbol{\lambda}, \boldsymbol{\theta}, Z_{\text{obs}}) p(\{X_1^{(k)}\}_{k=1}^K, \{X_2^{(k)}\}_{k=1}^K | \boldsymbol{\theta}, \boldsymbol{\phi}, \boldsymbol{\psi}) p(\boldsymbol{\theta}) p(\boldsymbol{\phi}) p(\boldsymbol{\psi}) p(\boldsymbol{\xi}).$ 

### CDFS: Pairwise marginal posterior distributions



Stampoulis (2018); Udaltsova (2014); Wang et al. (2022); Wright et al. (2015)