#### Volatility derivatives in (rough) forward variance models

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#### Reminders on forward variances

• Forward variance  $V_t^T$  are fair strikes of variance swaps :

payoff of Var swap over  $[t, T] = \frac{1}{T - t} \sum_{t_i \in [t, T]}^{N} \left( \log(S_{t_{i+1}}) - \log(S_{t_i}) \right)^2 - V_t^T$ 

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• Take  $T_2 > T_1$ . By combining positions in var swaps over  $[t, T_2]$  and  $[t, T_1]$ , we construct the payoff

$$\frac{1}{T_2 - T_1} \sum_{t_i \in [T_1, T_2]} \left( \log(S_{t_{i+1}}) - \log(S_{t_i}) \right)^2 - V_t^{T_1, T_2}$$

where

$$V_t^{T_1,T_2} = \frac{(T_2-t)V_t^{T_2} - (T_1-t)V_t^{T_1}}{T_2 - T_1}$$

is the forward variance over  $[T_1, T_2]$ .

#### Forward variances can be traded

- By entering in the opposite positions in variance swaps at a date t' ≥ t, we remove the realized variance part.
- We materialize a position depending only on forward variances :

portfolio value at 
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Otherwise said

$$\operatorname{price}_t \left( V_{t'}^{\mathcal{T}_1, \mathcal{T}_2} - V_t^{\mathcal{T}_1, \mathcal{T}_2} \right) = 0, \qquad \forall \, t \leq t' \leq \mathcal{T}_1$$

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Forward variances can be traded at zero cost.

▶ Under a pricing measure, the  $(V_t^{T_1, T_2})_{0 \le t \le T_1}$  have to be martingales

## Instantaneous forward variance $\xi_t^T$

• Define instantaneous forward variance by

$$\xi_t^T = \frac{d}{dT} \left( (T - t) V_t^T \right), \qquad t < T$$

$$V_t^T = \frac{1}{T-t} \int_t^T \xi_t^u \, du, \qquad t < T$$

and

$$V_t^{T_1,T_2} = \frac{1}{T_2 - T_1} \int_{T_1}^{T_2} \xi_t^u \, du, \qquad t < T_1 < T_2.$$

• Note that if  $\Delta$  is small, then

$$V_t^{T,T+\Delta} \approx \xi_t^T$$

$$\xi_t^T = \xi_0^T \exp\left(\int_0^t \mathcal{K}(T-s) \cdot dW_s - \frac{1}{2} \int_0^t \mathcal{K}(T-s) \cdot \rho \, \mathcal{K}(T-s) ds\right) \qquad t \le T$$

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(ξ<sub>0</sub><sup>T</sup>)<sub>T≥0</sub> is the initial forward variance curve – a market parameter.
W is a Brownian motion in ℝ<sup>n</sup> with correlation matrix ρ, and

$$\int_0^t \mathcal{K}(T-s) \cdot dW_s = \sum_{i=1}^n \int_0^t \mathcal{K}_i(T-s) dW_s^i$$
$$\int_0^t \mathcal{K}(T-s) \cdot \rho \, \mathcal{K}(T-s) ds = \sum_{i,j=1}^n \int_0^t \mathcal{K}_i(T-s) \rho_{i,j} \mathcal{K}_j(T-s) ds$$

▶ Deterministic kernels 
$$K_i \in L^2_{loc}(\mathbb{R}_+, \mathbb{R}^*_+)$$
.

$$\xi_t^T = \xi_0^T \exp\left(\int_0^t \mathcal{K}(T-s) \cdot dW_s - \frac{1}{2} \int_0^t \mathcal{K}(T-s) \cdot \rho \, \mathcal{K}(T-s) ds\right) \qquad t \le T$$

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• For every T,  $(\xi_t^T)_{t \leq T}$  is the solution of the SDE

$$\xi_t^T = \xi_t^T K(T-t) \cdot dW_t, \qquad t \leq T$$

- Does not belong to the affine family.
- Interest for simulation/calibration : only Gaussian r.v. are involved.

Choice of kernels in practice :  $\tau \mapsto K(\tau)$  decreasing.

## Parametric examples (I)

• Bergomi's model [Bergomi 05], [Dupire 93] with n=1 factor  $K( au)=\omega \ e^{-k au}$ 

with  $\omega, k > 0$ .

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$$\begin{aligned} \xi_t^T &= \xi_0^T \, \mathcal{E}\left(\omega \, \int_0^t e^{-k(T-s)} dW_s\right) \\ &= \xi_0^T \, \mathcal{E}\left(\omega \, e^{-k(T-t)} \int_0^t e^{-k(t-s)} dW_s\right) \\ &= \xi_0^T \, \exp\left(K(T-t)X_t - \frac{1}{2} \int_0^t K(T-s)^2 ds\right) \end{aligned}$$

where X is the OU process  $dX_t = -k X_t + dW_t$ .

For every t,  $\xi_t^T = \Phi(T - t, X_t)$ : the forward variance curve  $\xi_t^T$  is a function of one single Markov factor X.

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- For every t,  $\xi_t^T = \Phi(T t, X_t)$ : the forward variance curve  $\xi_t^T$  is a function of one single Markov factor X.
- Bergomi's *n*-factor model [Bergomi 05] is the *n*-dim extension :

$$K_i(\tau) = \omega_i \, e^{-k_i \tau}$$

## Parametric examples (II)

• The rough Bergomi model of [Bayer, Friz, Gatheral 2016] :

$$K( au)=rac{\omega}{ au^{rac{1}{2}-H}}, \qquad H\in(0,1/2).$$

so that

$$\xi_t^T = \xi_0^T \exp\left(\omega \int_0^t \frac{1}{(T-s)^{\frac{1}{2}-H}} dW_s - \frac{1}{2}\omega^2 \int_0^t \frac{1}{(T-s)^{1-2H}} ds\right)$$

Do not have a low-dimensional Markovian representation of the curve

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 $\blacktriangleright$  Do not have a low-dimensional Markovian representation of the curve  $\mathcal{T}\mapsto (\xi_t^{\, T})_{T\geq t}$ 

• For the moment (in this presentation), nothing in this model is rough. For every *T*, the processes

$$(\xi_t^T)_{t \leq T}$$
 are martingales

## Constructing a consistent model for $S_t$

Reminders : in a general stochastic volatility model

$$dS_t = \mu \, S_t dt + \sigma_t \, S_t \, dW_t^{\text{hist}}$$

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- Realized variance can be replicated with the underlying + a log-contract
- Indeed, by Itô's formula applied to log(S)

$$\frac{1}{T-t} \langle \log S \rangle_{[t,T]} = \frac{1}{T-t} \int_t^T \sigma_u^2 du = \frac{2}{T-t} \left( -\log \frac{S_T}{S_t} + \int_t^T \frac{1}{S_u} dS_u \right)$$

Almost sure replication of  $\langle \log S \rangle_{[t,T]}$ 

This yields (taking interest rates to be zero)

$$V_t^{\mathsf{T}} = \operatorname{price}_t \left( \frac{1}{\mathsf{T} - t} \int_t^{\mathsf{T}} \sigma_u^2 du \right) = \operatorname{price}_t \left( -\frac{2}{\mathsf{T} - t} \log \frac{\mathsf{S}_{\mathsf{T}}}{\mathsf{S}_t} \right)$$

## A consistent model for $S_t$

Given instantaneous forward variances  $\xi_t^T$ 

The model

$$dS_t = S_t \sqrt{\xi_t^t} \, dZ_t$$

where Z is a Brownian motion, is consistent with the given  $\xi_t^T$ 

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In the sense : the price of the log-contract in this model is

$$\operatorname{price}_{t}\left(\frac{-2}{T-t}\log\frac{S_{T}}{S_{t}}\right) = \mathbb{E}\left[\frac{1}{T-t}\int_{t}^{T}\xi_{u}^{u}du\Big|\mathcal{F}_{t}\right]$$
$$= \frac{1}{T-t}\int_{t}^{T}\mathbb{E}\left[\xi_{u}^{u}\Big|\mathcal{F}_{t}\right]du = \frac{1}{T-t}\int_{t}^{T}\xi_{t}^{u}du$$

where  $\mathcal{F}_t = \mathcal{F}_t^{W,Z}$ 

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• Hedging of European options on S with underlying + forward variances

#### Rough Bergomi model, again

• To see what is rough in rough Bergomi

we have to look at the consistent model for S :

$$dS_t = S_t \sqrt{\xi_t^t} \, dZ_t$$

The instantaneous volatility  $\xi_t^t$  of S is rough because

$$\xi_t^t = \exp\left(\omega \, \mathbf{x}_t^t - \frac{1}{2}\omega^2 \int_0^t \frac{1}{(t-s)^{1-2H}} ds\right)$$

and

$$x_t^t = \int_0^t rac{1}{(t-s)^{rac{1}{2}-H}} \, dW_s$$

is a Volterra process which admits a  $\beta$ -Hölder modification for  $\beta < H$ 

• The VIX is the price of the log-contract with 30 days maturity written on the SP500 :

$$\operatorname{VIX}_{t} := \sqrt{\operatorname{mkt price}_{t}\left(-\frac{2}{\Delta}\log\frac{S_{t+\Delta}}{S_{t}}\right)} \qquad \text{where } \Delta = 30 \text{days}$$

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• The value of VIX is quoted by the Chicago Option Exchange, by static replication of the payoff log(S) :

$$\mathrm{VIX}_{t} = \sqrt{\frac{2}{\Delta} \left( \int_{0}^{S_{t}} \frac{1}{K^{2}} \mathrm{P}_{t}(t + \Delta, K) dK + \int_{S_{t}}^{\infty} \frac{1}{K^{2}} \mathrm{C}_{t}(t + \Delta, K) dK \right)}$$

where  $P_t(T, K)$  and  $C_t(T, K)$  are market prices of put and call options on S, observed at t.

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 Is VIX an implied volatility? Yes, it is precisely the implied volatility of the log-contract.

## History of VIX (2006-2011)



#### VIX in a stochastic volatility model

• In general, VIX and forward variances of variance swaps do not coincide

$$\operatorname{VIX}_{t}^{2} \neq V_{t}^{t+\Delta} = \frac{1}{\Delta} \int_{t}^{t+\Delta} \xi_{t}^{u} \, du$$

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• Within a stochastic volatility model, on the contrary

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- Consequence : in general, we will be able to calibrate a forward variance model  $(S_t, \xi_t)$  to at most 2 of the 3 different markets :
  - VIX market
  - SP500 options market
  - Variance swap market on SP500

#### Pricing of VIX derivatives at t = 0

The price at t = 0 of a VIX option with payoff  $\varphi$  is

$$\mathbb{E}\left[\varphi(\mathrm{VIX}_{\mathrm{T}})\right] = \mathbb{E}\left[\varphi\left(\sqrt{V_{T}^{\mathcal{T}+\Delta}}\right)\right] = \Psi(0,\xi_{0}^{\cdot})$$

where

$$\Psi(0, x^{\cdot}) = \mathbb{E}\left[\varphi\left(\left(\frac{1}{\Delta}\int_{T}^{T+\Delta} x^{u} e^{\int_{0}^{T} \mathcal{K}(u-s) \cdot dW_{s} - \frac{1}{2}h(0, T, u)} du\right)^{1/2}\right)\right]$$
  
and  $h(t, T, u) = \int_{t}^{T} \mathcal{K}(u-s) \cdot \rho \mathcal{K}(u-s) ds.$ 

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and  $h(t, T, u) = \int_t^T K(u-s) \cdot \rho K(u-s) ds$ .

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- If Markov repr (e.g. classical Bergomi),  $\int_0^T K(u-s) \cdot dW_s = K(u-T) X_T$
- Otherwise : finite point (u<sub>i</sub>)<sub>i=1,...,N</sub> quadrature formula + simulation of the correlated Gaussian vector

$$\left(\int_0^T K(u_1,s) \cdot dW_s, \ldots, \int_0^T K(u_N,s) \cdot dW_s\right)$$

 $\rightsquigarrow$  see A. Jacquier's talk for rates of convergence.

Denote

 $\hat{\sigma}(t,T)$ 

the at-the-money implied volatility of an option on the forward volatility  $\sqrt{V_t^T}$ .

Proposition (ATM implied volatility of forward volatility)

The following asymptotics hold : for every T

$$\hat{\sigma}(t,T) \xrightarrow[t\to 0]{} \hat{\sigma}(0,T) := \frac{1}{2\int_0^T \xi_0^u du} \sqrt{\int_0^T \xi_0^u K(u) \cdot \rho \int_0^T \xi_0^{u'} K(u') du'}$$

• By choosing the kernels K, we can reach a prescribed target behavior of  $\hat{\sigma}(0,T)$ 



▶ Black dots : target behavior for  $\hat{\sigma}(0, T)$ , as a function of T (months).

• Very well described by a power law  $\frac{1}{T^{\alpha}}$ ,  $\alpha \approx 0.4 - 0.5$ 

• Choice 1 : n = 1 power kernel  $K(u) = \frac{\omega}{u^{\frac{1}{2}-H}}$ Then, if  $u \mapsto \xi_0^u$  is constant,

$$\hat{\sigma}(0, T) = \frac{\text{const.}}{T^{\frac{1}{2}-H}}$$

which is exactly our target term-structure, when  $H \approx 0.1$ .

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• Choice 2 : n = 2 exponential kernels

$$K_i(u) = \omega e^{-k_i u}$$
 and  $d \langle W^1, W^2 \rangle_t = \rho dt$ 

with  $k_1 \ll 1$ ,  $k_2 \gg 1$ .

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- A model with fractional kernel reaches the target behavior with n = 1 factor and two parameters ω, H.
- ► A classical Bergomi model does this with n = 2 factors and four parameters  $k_1, k_2, \rho, \omega$ .

As mentioned by Antoine, in the class of models above

- The  $\xi_t^T$  are log-normal. Forward variances  $\frac{1}{\Delta} \int_T^{T+\Delta} \xi_T^u du$  are close to log-normal.
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- Incapability of generate a reasonable smile for VIX options.
- Inspired by [Bergomi 2008], we set

$$\xi_t^T = \xi_0^T f^T(t, x_t^T)$$

where  $x_t^T$  denotes our Gaussian factor

$$x_t^T = \int_0^t K(T-s) \cdot dW_s$$

and the  $f^{T}(\cdot, \cdot)$  are smooth functions to be determined.

• We need to impose some conditions on  $f^T$ :

- $f^T(t,x) \geq 0$
- Initial condition  $\xi_0^T \Rightarrow = f^T(0,0) = 1, \forall T$
- $(\xi_t^T)_{0 \le t \le T}$  needs to be martingale :

$$d\xi_t^{\mathsf{T}} = \left(\partial_t f^{\mathsf{T}}(t, x_t^{\mathsf{T}}) + \frac{1}{2} \mathsf{K} \cdot \rho \mathsf{K} \ \partial_{xx} f^{\mathsf{T}}(t, x_t^{\mathsf{T}})\right) dt + \partial_x f^{\mathsf{T}}(t, x_t^{\mathsf{T}}) dx_t^{\mathsf{T}}$$

Therefore, we require that the  $f^{T}(\cdot)$  solve the family of PDE

$$\partial_t f^{\mathsf{T}}(t,x) + \frac{1}{2} K(\mathsf{T}-t) \cdot \rho K(\mathsf{T}-t) \partial_{xx} f(t,x) = 0, \qquad \forall (t,x) \in [0,\mathsf{T}] \times \mathbb{R}.$$

A simple representation : any  $C^{1,2}([0, T) \times \mathbb{R})$  function  $f^T$  with exponential growth satisfying the PDE above can be written in terms of its terminal condition

$$f^{T}(t,x) = \mathbb{E}\left[f^{T}(T,x+\sqrt{h(t,T,T)}G)\right]$$

where G is a standard Gaussian random variable (and recall that  $h(t, T, T) = \int_t^T K(T - s) \cdot \rho K(T - s) ds$ ).

- Positive solutions  $f^{T}(\cdot, \cdot)$  are parametrized by positive final conditions  $f^{T}(T, \cdot)$
- We can generate several parametric families of solutions.

#### Parametric choice 1 : polynomials

• The terminal condition :

$$f^{T}(T, y) = a(T)y^{2} + b(T)y + c(T)$$

leads to a quadratic Gaussian model

$$\xi_t^T = \xi_0^T f^T(t, x_t^T) = \xi_0^T \left( a(T) \left[ (x_t^T)^2 - h(t, T) \right] + b(T) x_t^T + 1 \right)$$
  
where  $h(t, T) = \int_0^t \mathcal{K}(T - s) \cdot \rho \mathcal{K}(T - s) ds$ 

▶ We are free to choose a(T), b(T) s.t.  $1 - a(T)h(T, T) - \frac{b(T)^2}{4a(T)} \ge 0$  (positivity condition).

• Example : if b(T) = 0,  $\xi_t^T$  has a  $\chi^2$  distribution.

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- Example : if b(T) = 0,  $\xi_t^T$  has a  $\chi^2$  distribution.
- The more general terminal condition :

$$f^{T}(T,y) = \sum_{k=0}^{n} a_{k}^{T} y^{2k}$$

leads to polynomial functions  $x \mapsto f^{T}(t, x)$ .

### Parametric choice 2 : exponentials

• The terminal condition :

$$f^{T}(T,y) = \sum_{k=1}^{m} \gamma_{k} e^{\omega_{k} y}$$
 where  $\sum_{k=1}^{m} \gamma_{k} = 1, \quad m \in \mathbb{N}$ 

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▶ The class of models we started from corresponds to m = 1.

• With this choice, forward variances

$$\xi_t^T = f^T(t, x_t^T)$$

are sums of log-normals which can be made very different from a single log-normal

• We have expressions and numerical methods for VIX derivatives similar to the previous case (where m = 1).

## A simple version of the rough model where forward variances are not log-normal

**Rough\_model** (n = 1, m = 2) : n = 1 gaussian factor, and m = 2 basis functions

$$f^{T}(t,x) = (1 - \gamma^{T}) \exp\left(\omega_{1}^{T} x - \frac{1}{2}(\omega_{1}^{T})^{2} h(t,T)\right) + \gamma^{T} \exp\left(\omega_{2}^{T} x - \frac{1}{2}(\omega_{2}^{T})^{2} h(t,T)\right)$$
$$\xi_{t}^{T} = \xi_{0}^{T} f^{T}(t,x_{t}^{T})$$
$$x_{t}^{T} = \int_{0}^{t} K(T - s) dW_{s}, \qquad K(T - s) = \frac{1}{(T - s)^{\frac{1}{2} - H}}$$

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This model depends on the global parameter

Н

and on the four term-structure parameters

$$\xi_0^T, \ \gamma^T, \ \omega_1^T, \ \omega_2^T$$

which we can use to fit an initial term-structure of VIX Futures and the smiles of VIX options.

#### Calibration to VIX market (m = 2 exponential fcts)

VIX Futures (left) and VIX implied volatilities (right) on 22 Nov 2017, T = 20 Dec



H = 0.1 (fixed)  $\xi_0^u|_{T \le u \le T + \Delta} = 0.0145$   $\gamma = 0.689$   $\omega_1 = 2.074$   $\omega_2 = 0.215$ 

Non-parametric choices of f leading to exact calibration

are possible

### Conclusion & further directions

• The consistent model for the SP500 :

$$dS_t = S_t \sqrt{\xi_t^t} \, dZ_t$$

might be a good candidate for a joint calibration of VIX and SP500 options

See the talk of J. Guyon at QuantMinds conference 2018 (former Global Derivatives), taking place this week, for some considerations about the feasibility of this joint calibration.

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In summary :

- Volterra Gaussian processes offer a considerable flexibility in the modeling of forward variances.
- Using more general functions than single exponentials allows to accomodate smiles of options on VIX, while keeping the Gaussian framework.
- "Rough" power kernels inevitably make the pricing of VIX Futures & options less tractable.
- ► Still accessible via Monte-Carlo + variance reduction.

## Thank you for your attention

Instead of instantaneous forward variances  $\xi_t^T$ , we can apply the framework above to model VIX Futures  $(\text{FVIX}_t^i)_{t < T_i}$  directly :

$$\text{FVIX}_{t}^{i} = \text{FVIX}_{0}^{i} f^{i}(t, x_{t}^{T_{i}}) \qquad T_{i} = \text{VIX maturities}$$

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Instead of instantaneous forward variances  $\xi_t^T$ , we can apply the framework above to model VIX Futures  $(\text{FVIX}_t^i)_{t \leq T_i}$  directly :

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- This opens the way to **non-parametric choices** of  $f^i$ :
  - ▶ VIX option prices imply a distribution  $\mathbb{P}^{mkt}(\text{FVIX}_{T_i}^i \leq K)$
  - Which we can exactly fit with the distribution of

$$\mathrm{FVIX}_{T_i}^i = \mathrm{FVIX}_0^i f^i (T_i, x_{T_i}^{T_i})$$

by choosing a monotone terminal function  $f^i(T_i, \cdot)$  (and using the fact that  $x_{T_i}^{T_i}$  is Gaussian).